

# Solutions 4

## Ch 3 Problems

3.12 Let  $E_1, E_2, E_3$  denote events that student passes 1st, 2nd, and 3rd exams

$$\text{So } P(E_1) = .9 \quad P(E_2|E_1) = .8 \quad P(E_3|E_1E_2) = .7$$

$$\text{Passes 1st \& 2nd: } P(E_2E_1) = P(E_2|E_1)P(E_1) = (.8)(.9) = .72$$

$$(a) \text{ Passes All } P(E_3E_2E_1) = P(E_3|E_1E_2)P(E_1E_2) = (.7)(.72) = .504$$

$$(b) \text{ P(failed second | did not pass all)} \\ = P(E_1E_2^c | (E_1E_2E_3)^c) = \frac{P((E_1E_2E_3)^c | E_1E_2^c)P(E_1E_2^c)}{P((E_1E_2E_3)^c)}$$

$$P((E_1E_2E_3)^c | E_1E_2^c) = 1$$

(if she fails second exam, she cannot pass all 3)

$$P((E_1E_2E_3)^c) = 1 - P(E_1E_2E_3) = 1 - .504 = .496$$

$$P(E_1E_2^c) = P(E_1) - P(E_1E_2) \quad \text{since } E_1 = E_1E_2 \cup E_1E_2^c \\ \text{mutually exclusive} \\ = .9 - .72 = .18$$

$$\text{So Answer} = \frac{(1)(.18)}{.496} = .362904$$

3.22 Roll red, blue, yellow dice  
 If  $R, B, Y$  are numbers appearing on the dice,  
 we ask for the probability of the event that  
 $B < Y < R$ .

(a)  $P(\text{no two dice on same \#})$

Sample space has  $6^3$  equally likely outcomes  
 There are  $6 \cdot 5 \cdot 4$  outcomes in which no number is repeated  
 $\begin{matrix} \nearrow & \uparrow & \uparrow \\ B & Y & R \end{matrix}$

$$\text{so } P(B \neq Y \text{ and } Y \neq R \text{ and } B \neq R) = \frac{6 \cdot 5 \cdot 4}{6^3} = \frac{10}{18}$$

(b) Given no two dice land on same #, what is probability  
 that  $B < Y < R$

The reduced sample space consists of  $6 \cdot 5 \cdot 4$  outcomes, all  
 equally likely. There are  $\binom{6}{3} = \frac{6 \cdot 5 \cdot 4}{3 \cdot 2 \cdot 1}$  possible

sets of numbers that can appear, and each set of  
 numbers can be distributed across the Blue, Yellow, Red  
 dice in  $3 \cdot 2 \cdot 1 = 6$  ways. For a given set of 3 numbers  
 there is only one of these 6 ways that satisfies  
 $B < Y < R$ , so

$$P(B < Y < R | \text{no two same}) = \frac{\left(\frac{6 \cdot 5 \cdot 4}{3 \cdot 2 \cdot 1}\right)}{6 \cdot 5 \cdot 4} = \frac{1}{3 \cdot 2 \cdot 1} = \frac{1}{6}$$

$$(c) P(B < Y < R) = P(B < Y < R | \text{no two same}) P(\text{no two same}) = \frac{10}{18} \cdot \frac{1}{6} = \frac{10}{108}$$

3.24

2 balls are either gold or black w/  $\frac{1}{2}$  probability, and independently from each other, placed in urn

4 outcomes each with  $\frac{1}{4}$  probability:  $\begin{cases} GG \\ GB \\ BG \\ BB \end{cases}$

(a) We know at least one ball is Gold  $F = \{GG, GB, BG\}$   
 Both gold =  $\{GG\}$   
 $P(\text{Both gold} \mid \text{at least one gold}) = P(\{GG\} \mid \{GG, GB, BG\})$

$$= \frac{P(\{GG\})}{P(\{GG, GB, BG\})} = \frac{\frac{1}{4}}{\frac{3}{4}} = \frac{1}{3}$$

(b) A gold ball falls out of urn, so the probability of the other being gold is  $\frac{1}{2}$ .

$$P(\text{both gold} \mid \text{discovered ball is gold}) = \frac{1}{2}$$

This is an apparent paradox, since it seems in both cases that we know one of the balls is gold. But in the second case, we also have to take into account the chance that the other ball could have fallen out.

So in (b) there are actually 8 possible outcomes

GG	GG	O = which one falls out.
GB	GB	
BG	BG	
BB	BB	

So we see that  $\Rightarrow$  exactly half the cases where a gold ball falls out, the other ball is gold.

3.37 Gambler has fair coin ( $P(H) = \frac{1}{2} = P(T)$ )  
and a 2-headed coin ( $P(H) = 1$ ,  $P(T) = 0$ )

(a) one coin is selected at random, flipped, comes up heads  
what is probability the coin is fair.

$$P(\text{fair} | H) = \frac{P(H | \text{fair}) P(\text{fair})}{P(H | \text{fair}) P(\text{fair}) + P(H | \text{2-headed}) P(\text{2-headed})}$$

$P(\text{fair}) = P(\text{2-headed}) = \frac{1}{2}$  since coin selected randomly

$$P(\text{fair} | H) = \frac{(\frac{1}{2})(\frac{1}{2})}{(\frac{1}{2})(\frac{1}{2}) + (1)(\frac{1}{2})} = \frac{\frac{1}{4}}{\frac{1}{4} + \frac{1}{2}} = \frac{\frac{1}{4}}{\frac{3}{4}} = \frac{1}{3}$$

(b) coin flipped again, heads again, what prob. coin is fair.

$$P(\text{fair} | HH) = \frac{P(HH | \text{fair}) P(\text{fair})}{P(HH | \text{fair}) P(\text{fair}) + P(HH | \text{2-headed}) P(\text{2-headed})}$$

$$= \frac{(\frac{1}{4})(\frac{1}{2})}{(\frac{1}{4})(\frac{1}{2}) + (1)(\frac{1}{2})} = \frac{\frac{1}{8}}{\frac{1}{8} + \frac{1}{2}} = \frac{\frac{1}{8}}{\frac{5}{8}} = \frac{1}{5}$$

(c) coin flipped, comes up tails what is probability coin is fair  
Obviously the coin must be fair:

$$P(\text{fair} | HHT) = \frac{P(HHT | \text{fair}) P(\text{fair})}{P(HHT | \text{fair}) P(\text{fair}) + P(HHT | \text{2-headed}) P(\text{2-headed})}$$

Since  $P(\text{HHT} | 2\text{-headed}) = 0$ , get

$$P(\text{fair} | \text{HHT}) = 1, \text{ as expected.}$$

3.47

Urn w/ 5 white, 10 black.

Fair die is rolled, that number of balls is drawn

$P(\text{all white?})$

Let  $R_i$  ( $i=1,2,3,4,5,6$ ) denote the event that the die rolls the number  $i$

$W$  = event that all balls drawn are white

$$P(W) = \sum_{i=1}^6 P(W | R_i) P(R_i), \quad P(R_i) = \frac{1}{6}$$

$$P(W | R_i) = \text{prob. of getting } i \text{ white when drawing } i = \frac{\binom{5}{i}}{\binom{15}{i}}$$

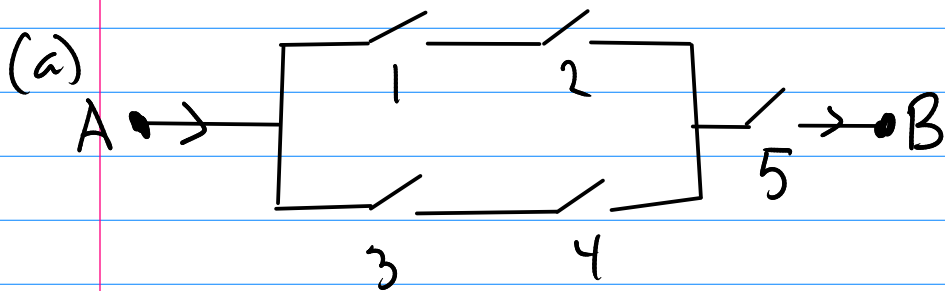
$$P(W) = \frac{\binom{5}{1}}{\binom{15}{1}} \frac{1}{6} + \frac{\binom{5}{2}}{\binom{15}{2}} \frac{1}{6} + \frac{\binom{5}{3}}{\binom{15}{3}} \frac{1}{6} + \frac{\binom{5}{4}}{\binom{15}{4}} \frac{1}{6} + \frac{\binom{5}{5}}{\binom{15}{5}} \frac{1}{6} + \frac{\binom{5}{6}}{\binom{15}{6}} \frac{1}{6} = 0$$

$$= \frac{1}{6} \left[ \frac{5}{15} + \frac{5 \cdot 4}{15 \cdot 14} + \frac{5 \cdot 4 \cdot 3}{15 \cdot 14 \cdot 13} + \frac{5 \cdot 4 \cdot 3 \cdot 2}{15 \cdot 14 \cdot 13 \cdot 12} + \frac{5 \cdot 4 \cdot 3 \cdot 2 \cdot 1}{15 \cdot 14 \cdot 13 \cdot 12 \cdot 11} + 0 \right]$$

$$= \frac{1}{6} \cdot \frac{5}{11}$$

$$\begin{aligned}
 P(R_3 | W) &= \frac{P(W | R_3) P(R_3)}{P(W)} \\
 &= \left[ \frac{\binom{5}{3}}{\binom{15}{3}} \right] \cdot \frac{1}{6} / P(W) = \left( \frac{2}{91} \right) \cdot \frac{1}{6} / \left( \frac{5}{11} \right) \cdot \frac{1}{6} \\
 &= \frac{22}{455}
 \end{aligned}$$

3.66 Let  $E_i$  denote the event that  $i$ th relay closes. These events are independent and  $P(E_i) = p_i$ .



Current flows  $\Leftrightarrow$  (1 & 2 close or 3 & 4 close) and 5 closes

$$= (E_1 E_2 \cup E_3 E_4) E_5$$

$$P((E_1 E_2 \cup E_3 E_4) E_5) = P(E_1 E_2 \cup E_3 E_4) P(E_5) \text{ by independence}$$

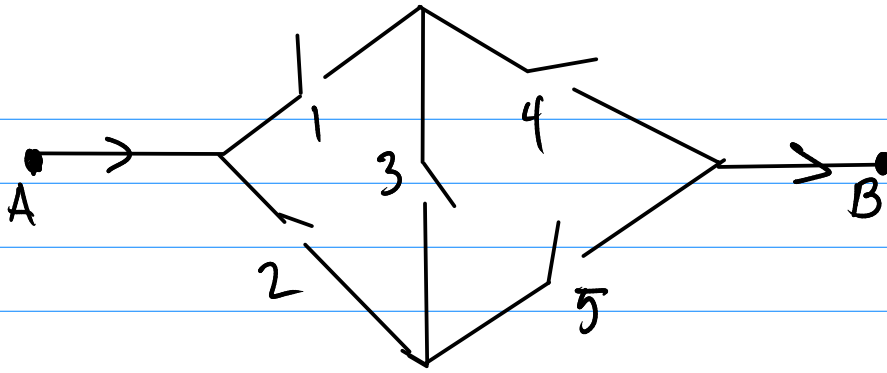
$$P(E_1 E_2 \cup E_3 E_4) \stackrel{\uparrow}{=} P(E_1 E_2) + P(E_3 E_4) - P(E_1 E_2 E_3 E_4)$$

by inclusion-exclusion

$$= p_1 p_2 + p_3 p_4 - p_1 p_2 p_3 p_4 \quad \text{by independence}$$

$$P(\text{current flows}) = [p_1 p_2 + p_3 p_4 - p_1 p_2 p_3 p_4] p_5$$

(b)



Condition on  $E_3$ :

If 3 closes, need (1 or 2) and (4 or 5)

$$P(\text{current flows} | E_3) = P((E_1 \cup E_2)(E_4 \cup E_5))$$

$$= P(E_1 \cup E_2) P(E_4 \cup E_5) \text{ by independence}$$

$$= [P(E_1) + P(E_2) - P(E_1 E_2)] [P(E_4) + P(E_5) - P(E_4 E_5)]$$

$$= [p_1 + p_2 - p_1 p_2] [p_4 + p_5 - p_4 p_5]$$

If 3 is open, need (1 and 4) or (2 and 5)

$$P(\text{current flows} | E_3^c) = P(E_1 E_4 \cup E_2 E_5)$$

$$= P(E_1 E_4) + P(E_2 E_5) - P(E_1 E_4 E_2 E_5)$$

$$= p_1 p_4 + p_2 p_5 - p_1 p_2 p_4 p_5$$

$$P(\text{current flows}) = P(\text{current flows} | E_3) P(E_3) + P(\text{current flows} | E_3^c) P(E_3^c)$$

$$= [p_1 + p_2 - p_1 p_2] [p_4 + p_5 - p_4 p_5] p_3 + [p_1 p_4 + p_2 p_5 - p_1 p_2 p_4 p_5] (1 - p_3)$$

3.78

A and B play series of games

Each game is independent,  $P(A \text{ wins}) = p$ 

$$P(B \text{ wins}) = 1 - p$$

Match ends when one player has won 2 more games than the other has won.

Note that a match always has an even # of games. A wins  $n$  times, B wins  $n+2$  times, or vice versa, so  $2n+2$  games are played

(a)  $P(4 \text{ games are played})$ 

length of match:	2	4
possible sequence of wins	AA	ABAA
	BB	ABBB
		BAAA
		BABB

Here, "A" denotes a win for A, B same.

$$\begin{aligned}
 P(\text{Match lasts 4 games}) &= P(ABAA) + P(ABBB) + P(BAAA) + P(BABB) \\
 &= p^3(1-p) + p(1-p)^3 + p^3(1-p) + p(1-p)^3 \\
 &= 2[p^3(1-p) + p(1-p)^3]
 \end{aligned}$$

(b)  $P(A \text{ wins match})$ 

The rules of this game are similar to the rules for a tie game of Tennis ("deuce"/"advantage" system)



We group the games into consecutive pairs:  
 Starting from a tie 4 things can happen

AA Tie  $\xrightarrow{A}$  advantage A  $\xrightarrow{A}$  A wins match

AB Tie  $\xrightarrow{A}$  advantage A  $\xrightarrow{B}$  tie

BA Tie  $\xrightarrow{B}$  advantage B  $\xrightarrow{A}$  tie

BB Tie  $\xrightarrow{B}$  advantage B  $\xrightarrow{B}$  B wins match.

So starting from tie, each sequence of two games results in either a tie or a win for one side.

So A wins  $\Leftrightarrow$  some sequence of BA or AB pairs, followed by AA

Let  $E_i =$  AB or BA on games  $2i-1$  and  $2i$

$F_i =$  AA on games  $2i-1$  and  $2i$

$$P(E_i) = \underset{A B}{p(1-p)} + \underset{B A}{(1-p)p} = 2p(1-p)$$

$$P(F_i) = p^2$$

These events are independent for different values of  $i$ .

A wins in  $2n+2$  games =  $E_1 E_2 \dots E_n F_{n+1}$

$$P(\text{A wins in } 2n+2 \text{ games}) = [2p(1-p)]^n p^2$$

$$\{A \text{ wins}\} = \bigcup_{n=0}^{\infty} \{A \text{ wins in } 2n+2 \text{ games}\}$$

$$P(A \text{ wins}) = \sum_{n=0}^{\infty} [2p(1-p)]^n p^2$$

$$= p^2 \sum_{n=0}^{\infty} [2p(1-p)]^n = p^2 \frac{1}{1-(2p(1-p))} = \frac{p^2}{1-2p+2p^2}$$

$$= \frac{p^2}{p^2 + (1-p)^2}$$


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3.4 Ball in one of  $n$  boxes  $P(\text{in } i\text{th box}) = P_i$   
 Search of  $i$ th box finds it with probability  $\alpha_i$

Let  $E_j =$  ball is in  $j$ th box

$S_i =$  search of  $i$ th box finds ball.

$$P(E_j | S_i^c) = \frac{P(S_i^c | E_j) P(E_j)}{\sum_{k=1}^n P(S_i^c | E_k) P(E_k)}$$

$$P(E_j) = P_j \quad P(S_i^c | E_j) = \begin{cases} 1 & \text{if } i \neq j \\ 1 - \alpha_i & \text{if } i = j \end{cases}$$

$$\text{Thus } \sum_{k=1}^n P(S_i^c | E_k) P(E_k) = \sum_{k \neq i} P_k + (1 - \alpha_i) P_i$$

$$= \left( \sum_{k=1}^n P_k \right) - \alpha_i P_i = 1 - \alpha_i P_i$$

$$\text{So } P(E_j | S_i^c) = \frac{P(S_i^c | E_j) P_j}{1 - \alpha_i P_i} = \begin{cases} \frac{P_j}{1 - \alpha_i P_i} & j \neq i \\ \frac{(1 - \alpha_i) P_i}{1 - \alpha_i P_i} & j = i \end{cases} \quad \square$$

3.6 Suppose  $E_1, E_2, E_3, \dots, E_n$  are independent

$$P(E_1 \cup E_2 \cup \dots \cup E_n) = P([E_1^c E_2^c \dots E_n^c]^c) \text{ by De Morgan}$$

$$= 1 - P(E_1^c E_2^c \dots E_n^c)$$

$$= 1 - \prod_{i=1}^n P(E_i^c) \quad \text{by independence}$$

$$= 1 - \prod_{i=1}^n [1 - P(E_i)]$$

3.10  $n$  people  $A_{i,j}$  = persons  $i$  and  $j$  have same birthday

$$P(A_{i,j}) = \frac{1}{365}$$

The events  $A_{i,j}$  &  $A_{r,s}$  for different pairs are independent

if  $\{i,j\} \cap \{r,s\} = \emptyset$  this is obvious

otherwise, we are looking at  $A_{i,j}$  and  $A_{j,k}$

$$P(A_{j,k} | A_{i,j}) = \text{Prob that } k \text{ has birthday on the particular day which } i \text{ and } j \text{ share} =$$

$$= \frac{1}{365} = P(A_{j,k})$$

But what about the three events  $A_{i,j}$ ,  $A_{j,k}$ ,  $A_{i,k}$

If these are to be independent, we need

$$P(A_{i,k} | A_{i,j} A_{j,k}) = \frac{1}{365}, \text{ But No!}$$

$i$  and  $j$  share birthday, and  $j$  and  $k$  share birthday, then  $i$  and  $k$  share birthday, necessarily

$$\text{so in fact } P(A_{i,k} | A_{i,j} A_{j,k}) = 1 \neq \frac{1}{365}$$

And these events are not independent all together.

3.16 Bernoulli trials:  $P_n = P(n \text{ trials result in even \# of successes})$

$E_n = n \text{ trials result in even \# of successes}$

$E_n^c = n \text{ trials result in odd \# of successes}$

$S_i = \text{success on the } i\text{th trial}$

condition on  $S_1$  (whether first trial is success)

$$P_n = P(E_n) = P(E_n | S_1) P(S_1) + P(E_n | S_1^c) P(S_1^c)$$

$$P(E_n | S_1) = P(E_{n-1}^c) \quad \text{b/c need odd \# of successes in next } n-1 \text{ trials}$$

$$P(E_n | S_1^c) = P(E_{n-1}) \quad \text{b/c need even \# of successes in } n-1 \text{ trials}$$

$$\text{So } P_n = (1 - P_{n-1})p + P_{n-1}(1-p)$$

$$\underline{\text{Claim:}} \quad P_n = \frac{1 + (1-2p)^n}{2}$$

Base case:

$$P_1 = P(S_1^c) = (1-p) = \frac{1 + (1-2p)^1}{2} \quad \checkmark$$

$$\text{Induction step: Suppose } P_{n-1} = \frac{1 + (1-2p)^{n-1}}{2}$$

$$P_n = p(1 - P_{n-1}) + (1-p)P_{n-1} = p + P_{n-1} - 2pP_{n-1}$$

$$= p + \frac{1 + (1-2p)^{n-1}}{2} - 2p \frac{1 + (1-2p)^{n-1}}{2}$$

$$= \frac{1}{2} \left[ 2p + 1 + (1-2p)^{n-1} - 2p(1 + (1-2p)^{n-1}) \right]$$

$$= \frac{1}{2} \left[ \cancel{2p} + 1 + (1-2p)^{n-1} - \cancel{2p} - 2p(1-2p)^{n-1} \right]$$

$$= \frac{1}{2} \left[ 1 + (1-2p)(1-2p)^{n-1} \right] = \frac{1}{2} \left[ 1 + (1-2p)^n \right] \quad \checkmark$$

QED.

3.18  $E_n =$  no run of 3 consecutive Heads in  $n$  coin flips

$$Q_n = P(E_n)$$

$$\text{Clearly } Q_0 = Q_1 = Q_2 = 1$$

Our sample space may be broken into 4 parts, depending on how the sequence starts

$$F_1 = \{ T \dots \dots \} \quad P(F_1) = 1/2$$

$$F_2 = \{ HT \dots \dots \} \quad P(F_2) = 1/4$$

$$F_3 = \{ HHT \dots \dots \} \quad P(F_3) = 1/8$$

$$F_4 = \{ HHH \dots \dots \} \quad P(F_4) = 1/8$$

$$Q_n = P(E_n | F_1) P(F_1) + P(E_n | F_2) P(F_2) + P(E_n | F_3) P(F_3) + P(E_n | F_4) P(F_4)$$

$$P(E_n | F_1) = P(\text{no 3 consecutive H on } n-1 \text{ trials}) = Q_{n-1}$$

$$P(E_n | F_2) = P(\text{no 3 consecutive H on } n-2 \text{ trials}) = Q_{n-2}$$

$$P(E_n | F_3) = Q_{n-3} \quad (\text{similarly})$$

$$P(E_n | F_4) = 0 \quad \text{since } F_3 \Rightarrow 3 \text{ consecutive heads}$$

$$\text{so } Q_n = \frac{1}{2} Q_{n-1} + \frac{1}{4} Q_{n-2} + \frac{1}{8} Q_{n-3}$$

$$Q_0 = Q_1 = Q_2 = 1$$

$$Q_3 = \frac{7}{8} \quad Q_4 = \frac{1}{2} \frac{7}{8} + \frac{1}{4} + \frac{1}{8} = \frac{13}{16}$$

$$Q_5 = \frac{1}{2} \frac{13}{16} + \frac{1}{4} \frac{7}{8} + \frac{1}{8} = \frac{24}{32}$$

$$Q_6 = \frac{1}{2} \frac{24}{32} + \frac{1}{4} \frac{13}{16} + \frac{1}{8} \frac{7}{8} = \frac{44}{64}$$

$$Q_7 = \frac{1}{2} \frac{44}{64} + \frac{1}{4} \frac{24}{32} + \frac{1}{8} \frac{13}{16} = \frac{81}{128}$$

$$Q_8 = \frac{1}{2} \frac{81}{128} + \frac{1}{4} \frac{44}{64} + \frac{1}{8} \frac{24}{32} = \frac{149}{256}$$