

Next HW

Problems 6.18, 6.20, 6.29, 6.30, 6.32, 6.38, 6.42

Theoretical

Exercises: 6.11, 6.14

Sums of independent Random variables

Suppose  $X$  and  $Y$  are independent random variables

Suppose  $X$  and  $Y$  are continuous

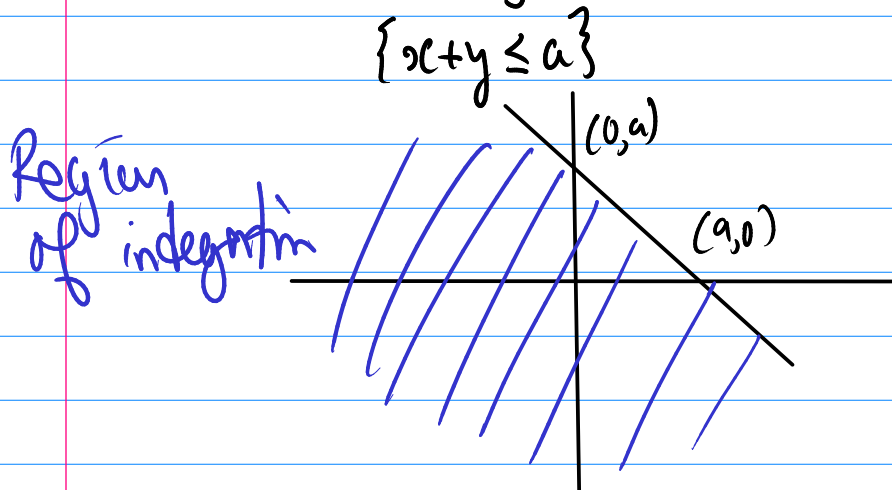
density functions  $f_X(x)$ ,  $f_Y(y)$ ,  $f_{X,Y}^{\text{joint}}(x,y)$

$X$  and  $Y$  independent  $\iff f_{X,Y}(x,y) = f_X(x) f_Y(y)$

Now consider  $X+Y$ : find distribution & density

$$F_{X+Y}(a) = P\{X+Y \leq a\}$$

$$= \iint_{\{x+y \leq a\}} f_{X,Y}(x,y) dx dy$$



$$x+y \leq a$$

$$x \leq a-y$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{a-y} f_{X,Y}(x,y) dx dy$$

So far, true  
for any R.V.s  $X, Y$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{a-y} f_X(x) f_Y(y) dx dy$$

using the fact  
that  $X$  and  $Y$   
are independent

$$= \int_{-\infty}^{\infty} \left[ \int_{-\infty}^{a-y} f_X(x) dx \right] f_Y(y) dy$$

$$= \int_{-\infty}^{\infty} F_X(a-y) f_Y(y) dy$$

$$\therefore F_{X+Y}(a) = \int_{-\infty}^{\infty} F_X(a-y) f_Y(y) dy$$

$$f_{X+Y}(a) = \frac{d}{da} F_{X+Y} = \int_{-\infty}^{\infty} \frac{d}{da} F_X(a-y) f_Y(y) dy$$

$$= \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$$

Definition If  $f$  and  $g$  are two functions  
The convolution of  $f$  and  $g$  is new function

$$(f * g)(x) = \int_{-\infty}^{\infty} f(x-y) g(y) dy$$

## Properties

$$(1) f * g = g * f$$

$$(2) (f * g)' = (f' * g) = (f * g')$$

To summarize if  $X$  and  $Y$  are independent R.V.s

$$\text{then } F_{X+Y} = F_X * f_Y = f_X * F_Y$$

$$f_{X+Y} = f_X * f_Y$$

Discrete case: integral replaced by sum

$$F_{X+Y}(a) = \sum_y F_X(a-y) P_Y(y)$$

$$= \sum_x P_X(x) F_Y(a-x)$$

$$P_{X+Y}(a) = \sum_y P_X(a-y) P_Y(y)$$

Example · Suppose  $X_1$  is normal with mean  $\mu_1$   
and variance  $\sigma_1^2$

· Suppose  $X_2$  is normal with mean  $\mu_2$   
and variance  $\sigma_2^2$

· Suppose  $X_1$  and  $X_2$  are independent

THEN  $X_1 + X_2$  is normal with mean  $\mu_1 + \mu_2$   
and variance  $\sigma^2 = \sigma_1^2 + \sigma_2^2$

ie.  $f_{X_1+X_2} = f_{X_1} * f_{X_2}$  we're saying

$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{(a-y-\mu_1)^2}{2\sigma_1^2}} \cdot \frac{1}{\sqrt{2\pi}\sigma_2} e^{-\frac{(y-\mu_2)^2}{2\sigma_2^2}} dy$$
$$= \frac{1}{\sqrt{2\pi}\sqrt{\sigma_1^2 + \sigma_2^2}} e^{-\frac{(a-\mu_1-\mu_2)^2}{2(\sigma_1^2 + \sigma_2^2)}}$$

Straightforward but lengthy computation

Example Basketball team  $P(\text{home win}) = .6$   
 $P(\text{away win}) = .5$

Play 41 home and 41 away.

$X = \# \text{ home wins} \sim \text{binomial with } n=41, p=.6$

$Y = \# \text{ away wins} \sim \text{binomial with } n=41, p=.5$

$X$  and  $Y$  are independent.

By DeMoivre-Laplace limit theorem

$$X \sim \text{normal w/ } \mu = np = 24.6 \\ \sigma^2 = np(1-p) = 9.84$$

$$Y \sim \text{normal w/ } \mu = np = 20.5 \\ \sigma^2 = np(1-p) = 10.25$$

Since "normal + normal = normal"

$$X+Y \sim \text{normal w/ } \mu = 24.6 + 20.5 = 45.1 \\ \sigma^2 = 20.09$$

$P(50 \text{ or more wins in the entire season})$

$$= P(50 \leq X+Y) = P(49.5 < X+Y)$$

$$= P\left(\frac{49.5 - 45.1}{\sqrt{20.09}} < Z\right) = 1 - \Phi\left(\frac{49.5 - 45.1}{\sqrt{20.09}}\right)$$

X and Y independent

X and Y are both exponential with rate  $\lambda$

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & x > 0 \\ 0 & x \leq 0 \end{cases} \quad f_Y(y) = \begin{cases} \lambda e^{-\lambda y} & y > 0 \\ 0 & y \leq 0 \end{cases}$$

X and Y are "independent identically distributed"

$$f_{X+Y}(a) = \int_{-\infty}^{\infty} f_X(a-y) f_Y(y) dy$$

in order for  $f_X(a-y) f_Y(y)$  to be nonzero

we need  $\begin{matrix} a-y > 0 \\ y > 0 \end{matrix} \Rightarrow y < a$

$$\rightarrow = \int_0^a \lambda e^{-\lambda(a-y)} \lambda e^{-\lambda y} dy$$

$$= \lambda^2 e^{-\lambda a} \int_0^a e^{+\lambda y} e^{-\lambda y} dy$$

$$= \lambda^2 e^{-\lambda a} a = \lambda e^{-\lambda a} \frac{(\lambda a)^1}{\Gamma(2)}$$

Observe that  $X+Y$  is a Gamma variable with rate  $\lambda$  and parameter  $\alpha=2$

$$\left( \begin{array}{l} \text{Gamma} \\ \lambda, \alpha \end{array} \quad f(x) = \lambda e^{-\lambda x} \frac{(\lambda x)^{\alpha-1}}{\Gamma(\alpha)} \right)$$

Observe that exponential RV. = Gamma random variable with  $\alpha=1$

in general - if  $X$  is Gamma w/  $\lambda, \alpha$

$Y$  is Gamma w/  $\lambda, \beta$

and  $X$  and  $Y$  are independent

then  $X+Y$  is Gamma w/  $\lambda, \alpha+\beta$

Conceptual explanation: if  $\alpha$  and  $\beta$  are integers

then  $X = \text{Gamma w/ } \lambda, \alpha = \text{time to wait for } \alpha \text{ events in Poisson process w/ rate } \lambda$

$Y = \text{Gamma w/ } \lambda, \beta = \text{time to wait for } \beta \text{ events}$

$X+Y = \text{time to wait for } \alpha+\beta \text{ events,}$

$= \text{Gamma w/ } \lambda, \alpha+\beta$