

Uniform and Normal Random Variables

Note: Theoretical Exercise 5.2, NOT 5.1
is on Homework

* Q-drop deadline is next Monday, April 2.

X is uniform on (α, β) if it has PDF

$$f(x) = \begin{cases} \frac{1}{\beta - \alpha} & \alpha < x < \beta \\ 0 & \text{otherwise} \end{cases}$$

Cumulative distribution function CDF

$$F(a) = \int_{-\infty}^a f(x) dx$$

$$\text{if } a \leq \alpha \quad \int_{-\infty}^a f(x) dx = \int_{-\infty}^a 0 dx = 0$$

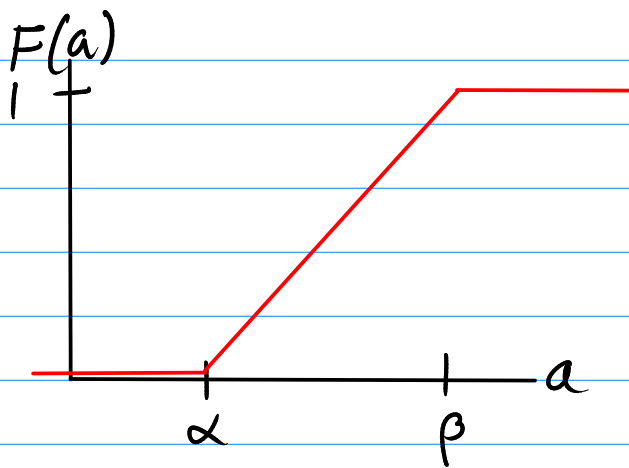
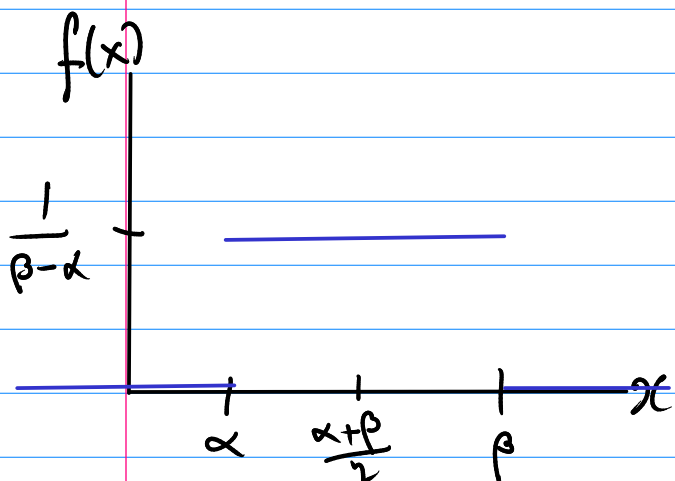
if $\alpha < a < \beta$

$$\int_{-\infty}^a f(x) dx = \int_{\alpha}^a f(x) dx = \int_{\alpha}^a \frac{1}{\beta - \alpha} dx$$

$$= \left[\frac{1}{\beta - \alpha} x \right]_{x=\alpha}^{x=a} = \frac{a - \alpha}{\beta - \alpha}$$

$$\text{if } \beta \leq a \quad \int_{-\infty}^a f(x) dx = \int_{\alpha}^{\beta} \frac{1}{\beta - \alpha} dx = \frac{\beta - \alpha}{\beta - \alpha} = 1$$

$$F(a) = \begin{cases} 0 & a \leq \alpha \\ \frac{a-\alpha}{\beta-\alpha} & \alpha < a < \beta \\ 1 & \beta \leq a \end{cases}$$



$$E[X] = \int_{-\infty}^{\infty} x f(x) dx = \int_{\alpha}^{\beta} x \frac{1}{\beta-\alpha} dx$$

$$= \frac{1}{\beta-\alpha} \left[\frac{x^2}{2} \right]_{\alpha}^{\beta} = \frac{1}{\beta-\alpha} \frac{1}{2} (\beta^2 - \alpha^2)$$

$$= \frac{1}{\cancel{\beta-\alpha}} \frac{1}{2} (\cancel{\beta-\alpha}) (\beta+\alpha) = \frac{1}{2} (\beta+\alpha)$$

= mid point of (α, β)

Variance: $E[X^2] = \int_{-\infty}^{\infty} x^2 f(x) dx = \int_{\alpha}^{\beta} x^2 \frac{1}{\beta-\alpha} dx$

$$= \frac{1}{\beta-\alpha} \left[\frac{x^3}{3} \right]_{\alpha}^{\beta} = \frac{1}{\beta-\alpha} \frac{1}{3} (\beta^3 - \alpha^3)$$

Fact $(\beta^3 - \alpha^3) = (\beta - \alpha)(\beta^2 + \alpha\beta + \alpha^2)$

$$E[X^2] = \frac{1}{3}(\beta^2 + \alpha\beta + \alpha^2)$$

$$\begin{aligned} \text{Var}(X) &= E[X^2] - (E[X])^2 \\ &= \frac{1}{3}(\beta^2 + \alpha\beta + \alpha^2) - \left(\frac{1}{2}(\beta + \alpha)\right)^2 \end{aligned}$$

$$= \frac{1}{3}(\beta^2 + \alpha\beta + \alpha^2) - \frac{1}{4}(\beta^2 + 2\alpha\beta + \alpha^2)$$

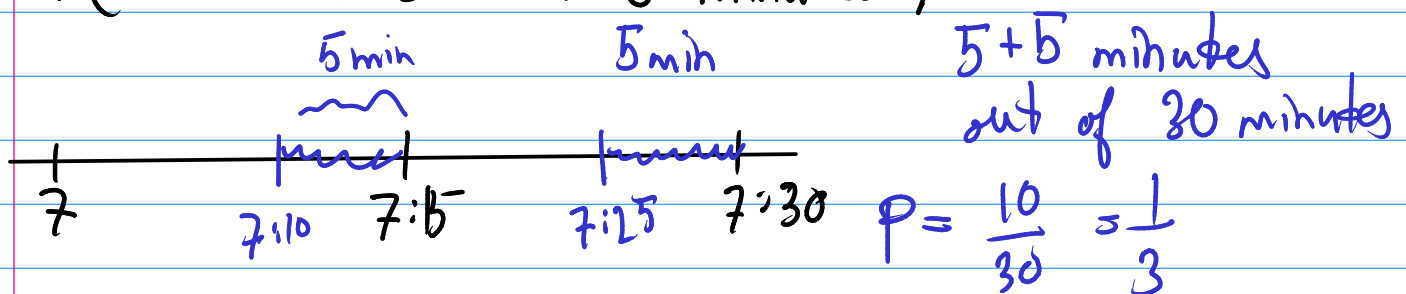
$$= \frac{1}{12}(\beta^2 - 2\alpha\beta + \alpha^2) = \frac{1}{12}(\beta - \alpha)^2$$

Observe $\beta - \alpha = \text{width of } (\alpha, \beta)$

Example Buses arrive at a stop at 7, 7:15, 7:30, ...

Passenger arrives at a time uniformly distributed in the interval (7:00, 7:30)

P(wait less than 5 minutes)



Normal Random Variable

a very "universal" distribution

→ approximates binomial distribution
(De Moivre - Laplace)

→ model errors in scientific observations (Gauss)

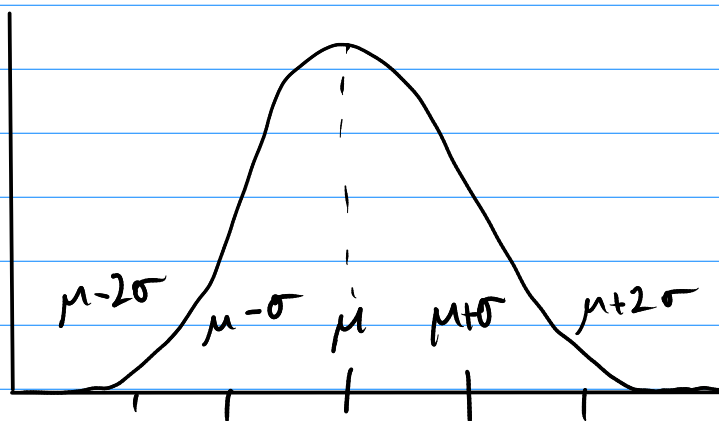
→ Models any sum of independent
identically distributed RVs with
finite mean and variance

(Central Limit Theorem)

Normal Random Variable with
mean μ & variance σ^2
(standard deviation σ)

has PDF

$$f(x) = \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \quad (-\infty < x < \infty)$$



Bell curve

Why is $\frac{1}{\sqrt{2\pi}\sigma} \int_{-\infty}^{\infty} e^{-(x-\mu)^2/2\sigma^2} dx = 1$?

$$\begin{aligned} &= \\ \text{sub } y &= \frac{x-\mu}{\sigma} \\ dy &= \frac{dx}{\sigma} \end{aligned} \quad \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-y^2/2} dy$$

Need to show $I = \int_{-\infty}^{\infty} e^{-y^2/2} dy = \sqrt{2\pi}$

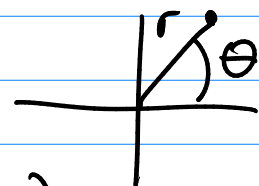
$$\begin{aligned} \text{Look at } I^2 &= \left(\int_{-\infty}^{\infty} e^{-y^2/2} dy \right) \left(\int_{-\infty}^{\infty} e^{-y^2/2} dy \right) \\ &= \left(\int_{-\infty}^{\infty} e^{-y^2/2} dy \right) \left(\int_{-\infty}^{\infty} e^{-x^2/2} dx \right) \end{aligned}$$

shuffle

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-y^2/2} e^{-x^2/2} dy dx$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-(x^2+y^2)/2} dy dx$$

change to polar coordinates



$$r^2 = x^2 + y^2 \quad \theta = \tan^{-1}\left(\frac{y}{x}\right)$$

$$dy dx = r dr d\theta$$

$$= \int_0^{2\pi} \int_0^{\infty} e^{-r^2/2} r \, dr \, d\theta$$

$$\int_0^{\infty} r e^{-r^2/2} \, dr$$

$$u = r^2/2 \quad du = r \, dr$$

$$= \int_0^{\infty} e^{-u} \, du = \left[-e^{-u} \right]_{u=0}^{u=\infty} = 0 - (-e^{-0}) = 1$$

$$I^2 = \int_0^{2\pi} 1 \, d\theta = 2\pi \Rightarrow I = \sqrt{2\pi}$$

Q: Why we just find the antiderivative of $e^{-y^2/2}$

$$\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-y^2/2} \, dy$$

Then by FTC, $\frac{d}{dx} \Phi = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$

Fact: $\Phi(x)$ cannot be expressed in terms of "elementary functions" (Liouville)

elementary
function

$$\frac{x^3 + x^2 + 2x + 5}{3x + \sqrt{x}}$$

algebraic function

$e^x, \ln x, \sin x, \arctan x$

elementary
transcendental
functions

$\Phi(x)$, $\Gamma(x)$, $\gamma(s)$, $J_n(x)$, ...
special functions

To deal with $\Phi(x)$, either

(a) lengthy computations by hand)

b) computer/calculator

c) Table of values (p. 201)

Standard normal Random variable

= Normal RV with mean 0 and variance 1

PDF $f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$

CDF $\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-y^2/2} dy$